

The Impact of Violence on Individual Risk Preferences

Evidence from a Natural Experiment

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Abstract

This study estimates the impact of Kenya's post-election violence on individual risk preferences. Because the crisis interrupted a longitudinal survey of more than five thousand Kenyan youth, this timing creates plausibly exogenous variation in exposure to civil conflict by the time of the survey. The study measures individual risk preferences using hypothetical lottery choice questions, which are validated by showing that they predict migration and

entrepreneurship in the cross-section. The results indicate that the post-election violence sharply increased individual risk aversion. Immediately after the crisis, the fraction of subjects who are classified as either risk neutral or risk loving dropped by roughly 26 percent. The findings remain robust to an IV estimation strategy that exploits random assignment of respondents to waves of surveying.

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The Impact of Violence on Individual Risk Preferences: Evidence from a Natural Experiment

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1 Introduction

Armed conflict is a source of untold human suffering. Since 1989, more than one million people have been killed in civil and interstate conflicts (Pettersson and Wallensteen 2015). The majority of these episodes are civil wars in low- and middle-income countries; because the scourge of war falls disproportionately on the poorest nations, armed conflict also perpetuates disparities in human and economic development among the living.

The short-term costs of civil conflict are obvious: in addition to the lives lost, war damages or destroys physical capital and deters investment. There is also evidence that war and violence limit the accumulation of human capital (Blattman and Annan 2010) and erode trust (Nunn and Wantchekon 2011). This has led some scholars to refer to civil war as “development in reverse” (Collier, Elliott, Hegre, Hoeffler, Reynal-Querol, and Sambanis 2003). Yet, though the short-term human and economic costs of conflict are indisputable, many conflict-affected countries — Rwanda and Uganda, for example — have experienced extremely rapid growth in the wake of civil war, and a number of recent papers have challenged the notion that conflict leads to slower growth and development over the long-term (cf. Miguel and Roland 2011). In fact, several studies have found that exposure to civil conflict increases political engagement (Bellows and Miguel 2009, Blattman 2009), enhances cooperation and pro-sociality (Voors, Nillesen, Verwimp, Bulte, Lensink, and Van Soest 2012, Bauer, Cassar, Chytilová, and Henrich 2013), and makes people more willing to bear profitable risks (Voors, Nillesen, Verwimp, Bulte, Lensink, and Van Soest 2012, Callen, Isaqzadeh, Long, and Sprenger 2014).

These studies share a common empirical approach. First, they take seriously the idea that exposure to conflict is endogenous, and employ a variety of strategies designed to isolate plausibly exogenous variation in victimization and involvement in violence. Second, given their focus on within-conflict variation in exposure and victimization, these papers empirically frame civilians who lived through civil war but were not victimized (or were less exposed to violence) as a comparison group. This strategy enhances the credibility of the estimated treatment effects, but has an obvious drawback: this approach can generate credible estimates of the marginal impact of greater conflict victimization or exposure, but cannot be used to assess the overall impact of conflict unless one

assumes that violence has no impact on the relatively less victimized. If everyone who lives through a period of conflict — regardless of their victim status — is affected, estimates of the marginal impact of greater conflict exposure may present a biased assessment of the overall social cost of violence.

In this paper, we estimate the impact of a specific episode of civil conflict, Kenya’s post-election crisis, on the risk preferences of a broad sample of young adults who lived through it. The post-election crisis was a months-long period of protests, rioting, and ethnic violence that began immediately after a disputed presidential election. The election, in which Raila Odinga challenged incumbent Mwai Kibaki, took place on December 27, 2007. Amidst allegations of electoral fraud by observers, and after three days of uncertainty following the national polls, the incumbent president was both declared the winner and sworn into office on December 30, 2007. Ethnic tensions rose, and rioting ensued. The following two months of civil conflict left more than a thousand people dead and hundreds of thousands more internally displaced. The crisis largely ended when, on February 28, 2008, the two candidates signed a power-sharing agreement.¹

We estimate the impact of Kenya’s post-election violence on individual risk preferences, which we measure using lottery choice questions embedded in a longitudinal survey. The Kenyan Life Panel Survey (hereafter KLPS2) is a survey of more than 5,000 young adults who were enrolled in rural primary schools in 1998. The second round of the survey was administered between August of 2007 and December of 2009. 1,180 respondents (23.3 percent) were interviewed prior to the crisis, while the remainder were surveyed after experiencing the period of civil conflict. Thus, Kenya’s post-election violence interacted with the timing of the survey to create a natural experiment in exposure to conflict.

We employ two complementary identification strategies to estimate the impact of the crisis on risk aversion. First, we estimate the impact of the crisis in straightforward linear and nonlinear frameworks, using several strategies to control for any time trends or seasonal shocks. Second, we exploit the fact that survey respondents were randomly assigned to one of two waves of interviews; Kenya’s election crisis interrupted the first wave of surveys, allowing us to instrument for conflict

¹ This description of Kenya’s 2007-2008 post-election crisis draws upon several sources: World Bank (2006), World Bank (2008), World Bank (2009), World Bank (2010a), World Bank (2010b), World Bank (2011), World Bank (2012), Van Praag (2010), Waki (2008), Hjort (forthcoming), McCrummen (2007), BBC (2008a), and BBC (2008b).

exposure (pre-survey) using the randomly-assigned survey waves. Both approaches yield similar results: we find that Kenya’s post-election crisis had a large and significant positive impact on individual risk aversion. Specifically, the crisis led to an 11 percentage point increase in the likelihood that a subject always chose the safest, lowest expected value alternative (i.e. lottery) available — this effect constitutes more than an 80 percent increase in the rate of extreme risk aversion. We also observe a 5.7 percentage point (or roughly 26 percent) decrease in the fraction of subjects who are classified as either risk neutral or risk loving. Such substantial impacts highlight an important channel through which civil conflict might affect growth and development: increased risk aversion might lead individuals in post-conflict settings to avoid high-risk, high-return activities (e.g. entrepreneurship) that contribute to economic growth.

The key strength of our study is that we are able to estimate the impact of civil conflict on the risk preferences of the general population, as opposed to the specific (marginal) effect of being victimized or more exposed to violence. Very few KLPS2 respondents were themselves victims of violence during the unrest: more than three quarters of respondents were temporarily deprived of basic necessities because it was not safe to visit markets or other public places, but less than 4 percent indicated that anyone in their household was physically assaulted during the crisis. Thus, KLPS2 respondents experienced the conflict, but were not, by and large, among the most impacted Kenyans; they therefore provide an important window into the impacts of civil conflicts on the preferences of the general population.

Our results differ from several recent studies. For example, Voors, Nillesen, Verwimp, Bulte, Lensink, and Van Soest (2012) find that greater exposure to violence leads to an increase in risk-seeking behavior, while Callen, Isaqzadeh, Long, and Sprenger (2014) find that priming subjects with recollections of violent events increases their risk tolerance in situations involving only uncertainty (but also increases their certainty premia). However, these differing results can be reconciled by the difference in estimands, for example, if those who narrowly escape being exposed to violence or victimized become more risk averse. Our aim in this paper is to estimate the overall impact of violence on the population as a whole; doing so, we may not find the same effect. Understanding these overall impacts is of critical importance as we seek to characterize the ways that conflict may change a country’s overall growth trajectory.

This paper contributes to several strands of literature. First, most obviously, we add to the evidence on the impacts of conflict. As discussed above, this literature has expanded rapidly in recent years as increasingly high-quality micro data from post-conflict settings has become available.² Second, we contribute to a growing body of evidence that individual preferences, one of the key determinants of individual behavior in all economic domains, are not as immutable as has long been assumed (cf. Stigler and Becker 1977); instead, individual preferences appear to be shaped by life experiences. For example, Malmendier and Nagel (2011) and Fisman, Jakiela, and Kariv (2015) show that exposure to economic downturns makes people more risk averse and more efficiency-focused, respectively; while Eckel, El-Gamal, and Wilson (2009), Cameron and Shah (2013), and Hanaoka, Shigeoka, and Watanabe (2015) estimate the impact of natural disasters on individual risk preferences (and arrive at different conclusions). As discussed above, several papers (cf. Voors, Nillesen, Verwimp, Bulte, Lensink, and Van Soest 2012) have estimated the impact of conflict on the preferences of those most affected, but, to our knowledge, no work to date has estimated the impact of violence on the risk preferences of the general population.

Finally, our study contributes to the growing body of evidence documenting the validity and predictive power of laboratory and lab-style measures of individual preferences. Though some scholars have questioned whether individual decisions in choice experiments predicts behavior outside of the lab (cf. Levitt and List 2007, Voors, Turley, Kontoleon, Bulte, and List 2012), numerous studies document the explanatory power of experimental measures of risk, time, and social preferences. For example, Liu (2013) and Liu and Huang (2013) show that experimental measures of risk preferences predict the crop choice and investment decisions of Chinese farmers. In the domain of time preferences, Meier and Sprenger (2012) show that experimental measures of patience predict creditworthiness. Fisman, Jakiela, and Kariv (2014) show that experimental measures of equality-efficiency tradeoffs predict the voting behavior of adult Americans, while Fisman, Jakiela, Kariv, and Markovits (2015) show that the same measure predicts the post-graduation career choices of law students. Jing (2015) shows that experimental measures of fair-mindedness predict medical students' choices regarding field of specialization. To date, the majority of work linking choices

²Humphreys and Weinstein (2006), Bellows and Miguel (2009), and Blattman and Annan (2010) are prominent examples. See Blattman and Miguel (2010) for discussion.

in decision experiments to behavior outside the lab has used incentivized measures of individual preferences, but there are notable exceptions (cf. Ashraf, Karlan, and Yin 2006). Existing evidence suggests the use of incentives shifts individual responses toward greater risk aversion (cf. Camerer and Hogarth 1999, Holt and Laury 2002), leading many to question the broad applicability of hypothetical approaches to risk preference elicitation. We contribute to this literature by demonstrating that hypothetical measures of risk preferences, interpreted as an ordinal index of risk tolerance, predict real world behaviors in an internally consistent way. In particular, our measure of risk preferences is associated with real-world behaviors that involve risk: migration and entrepreneurship.

The rest of this paper is organized as follows. In Section 2, we describe the KLPS2 data collection effort and Kenya’s post-election crisis. In Section 3, we describe our measure of risk preferences and assess the extent to which our hypothetical lottery choice questions predict behaviors likely to depend on risk aversion. In Section 4, we explain our analytic approach and present our main results. Section 5 concludes.

2 Research Design

2.1 The Kenyan Life Panel Survey

We measure the risk preferences of a large and heterogeneous sample of young Kenyans by embedding a series of non-incentivized decision problems in the second round of the Kenyan Life Panel Survey (KLPS2), which was administered to more than 5,000 individuals who were enrolled in primary schools in Kenya’s Western Province in the late 1990s.³ KLPS2 was administered in person, through one-on-one interviews, between August of 2007 and December of 2009. The survey covers a broad range of topics including educational attainment, labor market and entrepreneurial activities, household composition and wealth, migration, and fertility.

Our subject pool includes 5,049 Kenyan youths aged 14 to 31. Summary statistics characterizing

³KLPS2 was a long-term follow-up of the Primary School Deworming Project (PSDP) that was implemented in Busia District, in Kenya’s Western Province, between 1998 and 2001. All KLPS2 respondents were enrolled in public primary schools in the project area in 1998. See Miguel and Kremer (2004) for discussion of the PSDP and Baird, Blattman, Casaburi, Evans, Hamory, Kremer, Miguel, Ozier, Siegel, Wafula, and Yeh (2007) for a description of the KLPS data collection effort.

our sample are reported in the Online Appendix. 49.0 percent of subjects are female. As of 2007, subjects had completed an average of 8.5 years of schooling; 69.6 percent had completed primary school, and 14.4 percent had completed secondary school. 23.2 percent of subjects were still enrolled in school at the time that they were surveyed.

2.2 Kenya's Post-Election Crisis

The 2008 post-election crisis was a period of violence and political instability following the Kenyan presidential election of December 27, 2007. The two leading candidates in the election were Raila Odinga (the son of Kenya's first vice president), and the incumbent president, Mwai Kibaki (himself a former vice president). For several months preceding the election, opinion polls placed the opposition candidate, Odinga, ahead of Kibaki (World Bank 2008, Munene and Otieno 2007). As election day neared, the polling gap between Odinga and Kibaki narrowed to what, in some polls, appeared to be a statistical dead heat (World Bank 2008, Agina 2007, Otieno 2007, Standard Team 2007). At the time of the elections, both international observers and members of the opposition made allegations of electoral irregularities; however, after a few days, Kibaki was declared the winner and sworn into office (World Bank 2008, Waki 2008, McCrummen 2007, BBC News 2007). Opposition supporters organized protests; rioting and violent clashes with police ensued, and the violence soon escalated and took on a strong ethnic dimension (World Bank 2008, World Bank 2010b, Van Praag 2010, Waki 2008, BBC News 2007). By the time Kibaki and Odinga signed a power-sharing agreement at the end of February, 2008, more than one thousand people had been killed, and hundreds of thousands had been internally displaced (World Bank 2008, World Bank 2009, World Bank 2010a, World Bank 2012, BBC 2008a, BBC 2008b).

As discussed above, all KLPS2 respondents were residing in Busia District, in Kenya's Western Province near the Ugandan border, in 1998; 73 percent were still living in Busia District at the time of the KLPS2 survey, and many others happened to be there during the post-election crisis (because they had returned to their family homes to celebrate the Christmas holiday or to vote). As a result, most were spared the worst of the post-election violence: though protests, riots, and assaults took place throughout Kenya, the most affected areas were Rift Valley Province and the more urban districts across the country. According to the official report of the Commission of Inquiry on Post

Election Violence, only 98 of the 1,133 conflict deaths occurred in Western Province (Waki 2008). The majority of conflict deaths in Western Province occurred in districts that bordered Rift Valley Province; only 9 deaths were documented in Busia District (Waki 2008). In addition, most KLPS2 respondents (90.0 percent) are members of the Luhya ethnic group, the majority ethnic group in Western Province; as such, they were somewhat less likely to be singled out as clearly aligned with either the incumbent Mwai Kibaki (a member of a different ethnic group) or the opposition candidate Raila Odinga (a member of a third group).⁴⁵ Thus, though KLPS2 survey respondents lived through the crisis, both their physical locations and their ethnic identities helped to shield most of them from the worst of the violence.

This pattern is borne out in the data. The second wave of the survey included questions on exposure to violence during the post-election crisis; it documents the fact that very few respondents were, themselves, the victims of physical attacks.⁶ 76.3 percent of KLPS2 Wave 2 respondents reported that the crisis prevented them from going to local markets to obtain basic necessities, but only 3.7 percent indicated that someone in their household was physically assaulted during the crisis. Similarly low numbers of respondents had property stolen (5.8 percent) or burned (2.1 percent) during the crisis. Thus, though many of our subjects were directly affected by the crisis because businesses and markets were closed and people were confined to their homes during periods of rioting and insecurity, relatively few were, themselves, victims of violence. Our estimates should therefore be seen as both a lower bound on the impact of civil conflict on risk preferences, and a natural complement to studies which estimate the impact of violence on those individuals and households who were most victimized.

3 Measuring Risk Preferences

We elicit risk preferences by confronting survey respondents with a series of hypothetical decision problems; each decision was a choice between two or three lotteries involving two equally likely

⁴Western Province was an opposition stronghold, and Raila Odinga received 64.6 percent of the votes cast there in the presidential election; thus, individuals perceived as likely supporters of the incumbent were the most at risk.

⁵We thank the World Bank’s Kenya Country Office for guidance on wording.

⁶As discussed above, KLPS2 respondents were randomly assigned to one of two waves of surveying. The post-election violence interrupted the first wave of surveying. The additional survey were included only in Wave 2, so responses can be seen as representative of the entire respondent population.

potential payoffs. The sequence of decision problems was designed to start with choices that were extremely simple and to build slowly toward more complicated choice problems. For example, the first decision problem involved a choice between a degenerate lottery which paid 100 Kenyan shillings with certainty and a lottery involving a 50 percent chance of receiving 100 shillings and a 50 percent chance of receiving 120 shillings.⁷ In contrast, the final decision problem was a choice between three non-degenerate lotteries.

This sequencing of choice problems from least to most complex served two purposes. First, the gradual increase in complexity was intended to help address the concern that respondents might not exert sufficient cognitive effort (for example, to calculate expected payoffs) when facing non-incentivized choice problems.⁸ Respondents first considered options that could be evaluated with minimal cognitive effort, easing into the process of evaluating the expected utility of financial lotteries. Second, we wished to maximize the level of comprehension by subjects with low levels of numeracy. For this reason, we also limited each decision problem to a maximum of three lottery options, included only three easily understood probabilities (0, 0.5, and 1), and only considered lotteries over financial gains (rather than losses).⁹ Our experimental instructions do not assume any familiarity with probabilities, averages, or expected values; lotteries are explained in terms of payoffs and uncertain but equally likely events.

When administering our experiment, survey enumerators began by presenting two practice decision problems which introduced the structure of the lottery choice questions to the respondent.¹⁰ The first practice problem contained only degenerate lotteries — the respondent was asked to choose between 100 Kenyan shillings and 150 shillings; the second practice problem introduced the (non-degenerate) lottery concept in a setting with a clear “correct” answer: one lottery first-order

⁷The average US dollar value of 100 Kenyan shillings was 1.36 over the period during which the KLPS2 survey was in the field, from August 2007 to December 2009.

⁸There is some debate about the extent to which non-incentivized lottery choice questions correctly measure individual risk preferences. See Camerer (1995) for an overview, and Binswanger (1980) and Holt and Laury (2002) for seminal contributions. Existing evidence suggests that financial incentives induce greater risk aversion (Camerer and Hogarth 1999). Embedding non-incentivized decision problems in surveys has nonetheless proven successful in a number of field settings in developing countries, when conducting incentivized experiments at scale is not feasible (cf. Callen, Isaqzadeh, Long, and Sprenger 2014).

⁹Our risk preference elicitation method builds on those of Binswanger (1980), Eckel and Grossman (2008), Barr and Genicot (2008), Tanaka, Camerer, and Nguyen (2010), and Liu (2013), among others, but differs in its exclusive focus on extremely simple lottery structures and limited choice sets.

¹⁰Complete experimental instructions including images of the cards depicting lottery choice questions are included in the Online Appendix.

stochastically dominated the other, and both involved risk. Enumerators asked respondents to choose between the lotteries presented in the practice decision problems, and then followed a script which made sure that each respondent fully understood the nature of the choices they were facing.

After completing the practice decision problems, each subject made six choices between lotteries which differed in riskiness. Each choice was presented on a laminated card which depicted either two or three options. Like the practice problems, the first question (described above) provided a test of monotonicity, allowing us to identify subjects who are either not expected utility maximizers (cf. Andreoni and Sprenger 2010) or who simply could not grasp the nature of the choice problems. The second decision problem offered an extremely simple test of risk preferences, and one which should have been easily understood by almost all subjects: subjects were asked whether they preferred to receive 100 shillings with certainty, or a 50 percent chance of 400 shillings. The remaining four decision problems presented lottery choices of increasing complexity. Without any functional form assumptions (but assuming subjects are maximizing a well-defined utility function), these decisions can be used to classify subjects into risk preference categories: risk loving, risk neutral, moderately risk averse, and most risk averse (always choosing the lotteries with the lowest expected value and payoff spread). If we assume that consistent preferences can be represented by a utility function of the constant relative risk aversion (CRRA) form, the risk cards were calibrated to distinguish a wide range of coefficients in the neighborhood of 1.¹¹ The payouts for the lotteries in all the choice problems included in our experiment are described in the Online Appendix.

3.1 Individual Choices

Histograms of individual choices in all six decision problems are presented in the Online Appendix. Very few respondents (only 4.14 percent) indicated that they preferred the degenerate lottery which paid 100 Kenyan shillings with certainty to a non-degenerate lottery which paid either 100 or 120 Kenyan shillings, each with probability 0.5. We interpret this as evidence that most subjects understood the nature of the decision problems, at least those that involved relatively simple payoff

¹¹The CRRA utility function takes the form: $u(x) = x^{1-\rho}/(1-\rho)$ where ρ indicates the level of risk aversion. Higher values of ρ indicate greater risk aversion. When $\rho = 0$, the agent is risk neutral; the indifference curves represented by the CRRA utility function approach log utility as ρ approaches 1. Even if relative risk aversion is not constant, constant relative risk aversion is a reasonable approximation over the relatively small range of payoffs considered in our experiment. The CRRA form is also used in this setting in Jakiela and Ozier (forthcoming).

calculations.¹²

Using data from the other five decision problems allows us to assign respondents to distinct risk preference categories. We classify a subject as **risk neutral** if she always chose the lottery with the highest expected value. We classify a subject as **risk loving** if she made risk neutral choices in all except in the last decision problem, and then chose the lottery which paid 10 and 200 shillings with equal probability over the lottery which paid 70 and 160 shillings with equal probability. We classify those subjects who always opted for the lowest variance, lowest expected value lottery as the **most risk averse**. The three categories most risk averse, risk neutral, and risk loving account for 22.5, 2.4, and 15.7 percent of subjects, respectively.

Offering subjects a series of choices also allows us to test whether individual decisions are consistent with a CRRA utility representation. There are 162 possible combinations of responses to the last 5 lottery choice problems, only 10 of which are consistent with a CRRA utility representation. 44.6 percent of subjects made choices which were consistent with the maximization of a CRRA utility function. If subjects were choosing lotteries at random, only 6.2 percent would make CRRA-consistent choices. However, only 4 percent of subjects made consistent choices that were not classified as either most risk averse, risk neutral, and risk loving. Figure 1 presents the fraction of respondents falling into each of five different categories: most risk averse, risk neutral, risk loving, other types of consistent choices (suggesting intermediate levels of risk aversion), and inconsistent choices. The figure demonstrates that patterns of choices which are inconsistent with the maximization of a CRRA utility function occur substantially less frequently than we would expect if subjects were selecting lotteries at random, but that only the first three types of consistent

¹²We also asked enumerators to indicate whether they believed that respondents fully understood the lottery choice questions; these responses suggest that 99.6 percent of subjects fully comprehended the choices that they were making. There are, of course, several reasons that subjects who understood the decision problem might prefer a degenerate lottery to a stochastically dominant one involving risk. One possibility is that some respondents have preferences which are not monotonic. Several non-expected utility models of risk preferences suggest that individuals may prefer to avoid increases in payoff variance, even when the higher variance lottery stochastically dominates the lower variance alternative. For example, specifications of the non-linear probability weighting component of prospect theory typically assume sub-additivity (Kahneman and Tversky 1979); the models of risk attitudes proposed by Kösegi and Rabin (2006) and Andreoni and Sprenger (2010) are also consistent with a preference for the risk-free Option A over the stochastically dominant Option B. We interpret the overwhelming tendency to choose the stochastically dominant lottery as evidence against strongly non-monotonic preferences. One plausible interpretation of the observed choice patterns is that respondents have monotonic preferences which they implement with error (Hey and Orme 1994, Loomes 2005, Von Gaudecker, van Soest, and Wengström 2011, Choi, Kariv, Müller, and Silverman 2014).

behavior (extreme risk aversion, risk neutrality, and risk loving behavior) occur more frequently than random choice would suggest.

3.2 Instrument Validity

To address concerns about the validity of hypothetical measures of risk aversion, we test whether individual responses to our hypothetical lottery choice questions explain observed variation in behaviors that we might expect to be driven (at least in part) by risk attitudes. We focus on two such behaviors suggested by the literature: entrepreneurship (Schumpeter 1934, Schumpeter 1939, Skriabikova, Dohmen, and Kriechel 2014) and migration in search of employment (Jaeger, Dohmen, Falk, Huffman, Sunde, and Bonin 2010, Bryan, Chowdhury, and Mobarak 2014). Both of these actions can be seen through the lens of risk tolerance: for an unemployed or underemployed young adult residing in a rural area, operating one's own business and migrating are two high risk but potentially profitable strategies for improving one's long-term income prospects. Both behaviors are also relatively uncommon: only 12.4 percent of KLPS2 respondents were operating their own business at the time of the survey, and only 3.5 percent had ever moved for a job or in search of work.

To explore the association between these risky but potentially profitable activities and responses to our lottery choice questions, we construct a simple index of risk tolerance: a count of the number of times (out of five lottery choice questions) that a subject chose the riskiest and highest expected value lottery.¹³ This index has two important strengths. First, it obviates the need to impose functional form assumptions on the structure of risk preferences. In the decision problems included in our study, all lotteries involving risk have two equally likely outcomes, and the highest expected value lottery always has the highest variance and the lowest minimum payoff (i.e. the lowest payoff in the bad state),¹⁴ so any parametric restriction that permits an ordering of utility functions in

¹³Focusing on the number of risky choices allows us to include those subjects whose decisions were not consistent with a CRRA utility representation in our analysis. As discussed in Loomes (2005) and Choi, Kariv, Müller, and Silverman (2014), most subjects implement their choices with error — and we would expect relatively high levels of error in our sample because of the relatively low level of formal schooling attained by our subjects. Importantly, if the likelihood of making consistent choices is associated with risk aversion, omitting the inconsistent subjects could bias our results. However, all of our findings are robust to the omission of the inconsistent types from the analysis.

¹⁴The one exception to this is that the final lottery choice question included one alternative designed to distinguish between risk loving and risk neutral types. In this choice problem, we classify both the highest expected value lottery and the highest variance lottery as risky, though we obtain similar results if we do not classify the highest expected

terms of risk aversion would predict that relatively more risk averse individuals will be less likely to choose the highest variance lotteries. Second, constructing a measure of risk aversion that does not rely on functional form assumptions allows us to use data from all subjects, including those whose choices were not perfectly consistent with a specific utility representation.

We estimate probit regressions of indicators for operating one’s own business and having ever migrated seeking work on our index of risk aversion: the number of times (out of five) that a respondent chose the riskiest (highest expected value) lottery. Results are reported in Table 1. In all specifications, the number of times a respondent opted for the riskiest lottery is significantly associated with the likelihood of operating one’s own business or migrating for work. After controlling for gender, education level, age, the month in which the survey took place, and the survey enumerator who administered the lottery choice module, an additional risky choice is associated with a 0.5 percentage point (4.3 percent) increase in the likelihood of operating a business and a 0.2 percentage point (6.7 percent) increase in the likelihood of having ever migrated for work. Given the many factors underlying occupational choices, we interpret these robust associations as clear evidence that our lottery choice measure does predict outcomes associated with risk aversion in the cross-section.

4 Analysis

4.1 Empirical Approach

We exploit the fact that Kenya’s post-election crisis occurred in the middle of the KLPS2 data collection effort to estimate the impact of the violence on risk preferences. The KLPS2 survey was launched in August of 2007, and 1,180 respondents were surveyed in 2007 prior to the elections. The crisis led to a two-month suspension of survey activities, which resumed in March of 2008 and continued through the end of 2009. We employ two complementary identification strategies. First, we estimate the impact of the crisis in a straightforward linear framework, controlling for month-of-the-year fixed effects. In these specifications, we also include enumerator fixed effects and controls for sociodemographic characteristics such as respondent age, educational attainment, value lottery as a risky alternative.

and marital status that we expect to change over time.¹⁵ Thus, we estimate OLS regressions of the form:

$$Y_{ijm} = \alpha + \beta Post_i + \delta X_i + \eta_j + \lambda_m + \varepsilon_{ijm} \quad (1)$$

where Y_{ijm} denotes the risk aversion index of respondent i surveyed by enumerator j in month m , $Post_i$ is an indicator variable equal to one if respondent i was surveyed after the post-election crisis, η_j is an enumerator fixed effect, λ_m is a month (of the year) fixed effect, and ε_{ijm} is a conditionally mean-zero error term. We also report OLS specifications that replace the month fixed effects with separate linear time trends for the pre- and post- crisis periods.¹⁶ As a robustness check, we also take a nonlinear approach, and report analogous ordered logit specifications which account for the ordered nature of our outcome variable, the number of times a respondent chose the riskiest, highest expected value lottery.

We also report complementary instrumental variables estimates of the impact of the post-election crisis on individual risk preferences, exploiting the fact that KLPS2 respondents were randomly assigned to one of two waves of surveying. Wave 1 began in August of 2007 and continued through November 2008, while Wave 2 started in November 2008 and concluded in December of 2009. As discussed above, Kenya’s post-election crisis occurred approximately halfway through the first wave of surveys: 1,180 Wave 1 surveys were completed before the crisis, and 1,289 were completed afterward. This enables us to use random assignment to Wave 2 as an instrument for being surveyed after the crisis. Within each wave, participant characteristics may be associated

¹⁵With the exception of age, these characteristics might also be impacted by the crisis. To address concerns about the potential for a bad controls problem, we also report specifications that omit all sociodemographic controls.

¹⁶While it is tempting to view our research design as analogous to a regression discontinuity approach, our natural experiment does not create a valid discontinuity. First, because the post-election crisis halted surveying for more than two months, sociodemographic characteristics such as age do jump discontinuously at the moment of the crisis. Second, because the crisis lasted months, the limit-based notion of regression discontinuity is not directly applicable to this setting; a discontinuity would involve choosing a single point in time during the crisis, and at any such point, data from this survey would not be immediately adjacent on both sides. Third, there is some evidence that neither the onset nor the conclusion of the crisis were, in fact, instantaneous and completely unexpected. Though opposition candidate Raila Odinga held a convincing lead in opinion polls several months before the election, incumbent Mwai Kibabki caught up with Odinga in the month prior to the polling date, potentially raising fears that one of the candidates might try to manipulate a close vote. The official report Commission of Inquiry also documents several minor clashes in Rift Valley in late November and early December, most of which appeared be connected to attempts by political candidates to rally support by playing off of underlying ethnic tensions and disagreements over land ownership (Waki 2008). There was also a period of uncertainty in the wake of the crisis, between February 28, 2008, when the power-sharing agreement was signed, and April 12 when — after sometimes tense negotiations between the two political parties — Raila Odinga was sworn in as Prime Minister and the new coalition government officially took power.

with how quickly an individual was surveyed — for example, subjects who were still residing with their parents in their home villages might have been surveyed earlier because they were easier to locate. However, random assignment to survey wave creates exogenous variation in exposure to civil conflict prior to the survey. Thus, our 2SLS regressions are of the form:

$$Y_{ijm} = \alpha_1 + \beta_1 Post_i + \delta_1 X_i + \eta_{1j} + \lambda_{1m} + \zeta_{1ijm} \quad (2)$$

$$Post_i = \alpha_2 + \gamma_2 Wave2_i + \delta_2 X_i + \eta_{2j} + \lambda_{2m} + \zeta_{2ijm} \quad (3)$$

where all variables are as before, β_1 is the coefficient of interest, $Wave2_i$ is an indicator variable for the randomized survey timing instrument, and both ζ_{1ijm} and ζ_{2ijm} are conditionally mean-zero error terms. The variable $Wave2_i$ takes the value 1 for individuals assigned to be surveyed in Wave 2 (which is entirely post-crisis), and 0 for those assigned to be surveyed in Wave 1 (which includes both pre- and post-crisis observations).

4.2 Results

We begin with a graphical presentation of our key result: Figure 2 plots the average proportion of hypothetical questions in which the respondent chose the riskiest, highest expected value lottery, as a function of the month and year in which the survey was administered. The figure highlights the marked drop in the tendency to choose risky lotteries after the violence: respondents chose the riskiest lottery an average of 2.6 times in 5 decision problems prior to the post-election violence; this figure drops to 1.9 times out of 5 after the violence. In fact, the average rate of risky lottery choice is higher in all 5 pre-crisis months than in any of the following 18 months.

We now proceed to estimate the impact of exposure to the post-election crisis on the risk preferences by estimating Equation 1. Our dependent variable is the number of times a respondent chose the riskiest, highest expected value lottery (over the course of 5 decision problems). Results are reported in Table 2. In Column 1, we include only the indicator for being surveyed after the post-election crisis. In Column 2, we control for gender, age, education level, and marital status at the time of the survey. We add survey enumerator fixed effects in Column 3, and include controls for both the survey enumerator and respondent demographic characteristics in Column 4. Columns

5 through 8 replicate 1 through 4 but include fixed effects for the month of the year in which the survey took place.

In all eight OLS specifications reported in Table 2, the post-election coefficient is negative and significant at at least the 99 percent confidence level, suggesting that experiencing the crisis lowered the number of risky choices a respondent made by between 0.561 and 0.798 choices. In the pre-crisis period, the average number of risky choices was only 2.566 out of five, so the change observed after the crisis represents a dramatic drop in the willingness to bear profitable risk. Adding controls for gender, education level, age, and marital status has almost no impact on the estimated coefficient. In Table 3, we estimate analogous ordered logit specifications which account for the discrete but ordered nature of the outcome, the number of risky choices. Again we find that the indicator for being surveyed after the post-election violence is negative and significant (at at least the 99 percent confidence level) in all specifications. In Table 4, we take an alternative approach to controlling for time effects by including separate linear time trends in the pre- and post- crisis period. Columns 1 through 4 report OLS specifications, and Columns 5 through 8 report ordered logit results. Again, the indicator for being surveyed in the post-crisis period is negative and significant in all specifications.

As an additional set of robustness checks, we consider alternative formulations of the outcome variable in Table 5. In the first column, the outcome variable is an indicator for always choosing the highest variance and/or highest expected value lottery option (i.e. behaving in a manner consistent with risk neutral or risk-loving preferences); in the second column, the outcome is an indicator for always choosing the lowest variance alternative (i.e. behaving in a manner consistent with extreme risk aversion). We find that the post-election violence led to a 5.7 percentage point (26 percent) decrease in the likelihood of making risk neutral or risk loving choices, and an 11 percentage point (82 percent) increase in the probability of always choosing the lowest variance (and lowest expected return) lottery option. Next, we consider each decision problem in isolation. Importantly, this allows us to test whether the effect that we find is likely to be driven by a change in the preference for certainty, a finding that would be consistent with the results reported in Callen, Isaqzadeh, Long, and Sprenger (2014).¹⁷ We find that the post-election violence led to a significant decrease

¹⁷Callen, Isaqzadeh, Long, and Sprenger (2014) find that Afghan respondents primed to recollect fearful episodes

in the likelihood of choosing the highest variance, highest expected value lottery option in each of the five decision problems. Moreover, the effect is similar in magnitude across the four decision problems that include three lottery alternatives; the estimated effect of violence does not appear to be larger in decision problems that include a degenerate (i.e. risk-free) lottery as one of the alternatives. We interpret this as evidence that the impact of violence does not only operate through an increased preference for certainty.¹⁸ Thus, our main result is robust to a wide range of alternative specifications.

Next, we employ a complementary identification strategy, exploiting the fact that KLPS2 respondents were randomly assigned to one of two waves of surveying to generate instrumental variables (IV) estimates of the impact of the post-election violence on measured risk aversion.¹⁹ IV regression results are reported in Table 6. In Column 1, we include only controls for gender, age, education level, and marital status. Coefficient estimates indicate that the post-election crisis had a large and statistically significant impact on measured risk preferences, reducing the number of times (out of five) that respondents opted for the riskiest, highest expected value lottery by approximately one full question. Again, relative to a mean number of risky choices of about 2.5 in the pre-crisis period, this is an extremely large effect. In Columns 2 through 4 we include additional controls for the identity of the survey enumerator and the month in which the interview took place. Coefficient estimates are consistently negative and significant. Thus, our IV estimates are consistent with our un-instrumented results: Kenya’s post-election crisis appears to have made survey respondents substantially more risk averse.

display greater risk aversion when certainty is an option, but greater risk tolerance when no risk-free alternatives are available. They interpret this as evidence that subjects display a preference for certainty not consistent with expected utility maximization; their evidence suggests that violence impacts this preference for certainty.

¹⁸To further probe this issue, we examine the likelihood of choosing the *safest* (i.e. lowest variance, lowest return) lottery alternative in each of the five decision problems. Results are reported in the Online Appendix. We find no evidence that the shift away from high-risk lotteries is stronger in decision problems that included options with purely certain outcomes.

¹⁹Before proceeding to our IV analysis, we check whether the random assignment of respondents to survey waves generated groups that were comparable in terms of observable characteristics at the start of the survey. We focus on fixed traits and outcomes that were recorded in the survey for specific points in time. For example, the survey collects detailed information about school participation in each year between 1998 and the moment of the interview; this allows us to generate a variable indicating the number of years of schooling that a respondent had completed at the start of 2007, regardless of the year in which a respondent was actually surveyed. Results are reported in the Online Appendix. We find no statistically significant differences between the survey waves in terms of proportion female, age in 2007, proportion from the local-majority Luhya ethnic group, years of schooling completed by 2007, and the proportion that were married by 2007. Thus, the randomization appears to have succeeded in creating groups of respondents that are similar in terms of observable characteristics.

4.3 Potential Mechanisms

We interpret our results as evidence that civil conflict impacts individual risk preferences. One alternative explanation is that the post-election violence was an economic shock, and that young adults expected it to slow growth and reduce job opportunities; if those with higher expected permanent incomes are more willing to bear risk, such a shock might have impacted risk preferences through job prospects. In the Online Appendix, we show that the post-election crisis did not have a significant impact on the likelihood of being engaged in any off-farm income-generating activity. In fact, there is some evidence that the post-election violence led to an increase in the likelihood of working for a wage, though the effect is not significant in the IV specification. As expected given the estimated impact on risk preferences, we find that self-employment is significantly less common after the violence — highlighting a potential channel through which conflict may slow growth and development.

Voors, Nillesen, Verwimp, Bulte, Lensink, and Van Soest (2012) suggest another potential mechanism through which conflict might change preferences and behavior: the social context may change in the wake of violence. We find some evidence consistent with this hypothesis: after the violence, KLPS2 respondents displayed significantly less generalized trust (as measured through their responses to the standard question: “Generally speaking, do you think that most people can be trusted, or would they try to take advantage of you if they got the chance?”).²⁰ Violence may impact many aspects of society, and it is impossible to fully distinguish between mechanisms and other outcomes. Moreover, as others have noted, the willingness to trust and the willingness to take risks are, from a theoretical perspective, closely related (Karlan 2005, Schechter 2007). However, the evidence suggests that the post-election violence did not impact risk preferences purely through a labor market channel; instead, we find suggestive evidence that the violence changed the social context by altering both preferences and beliefs about others.

²⁰Results reported in the Online Appendix.

5 Conclusion

We measure the impact of Kenya’s post-election violence on individual risk preferences. We find that experiencing the post-election crisis appears to have increased measured risk aversion significantly. Point estimates suggest that exposure to the crisis decreased the likelihood of making risk neutral or risk loving choices by 5.7 percentage points (26 percent), and increase the likelihood of always choosing the lowest variance, lowest expected value lottery by 11 percentage points (82 percent). Our results are robust to a range of controls and the use of entirely distinct identification strategies. Thus, the evidence suggests that exposure to the post-election crisis led to a statistically and economically significant decrease in the willingness to take profitable risks.

In relation to existing studies of the determinants of individual preferences, our results corroborate a growing body of evidence that preferences are impacted by major life events such as conflict, disasters, and economic downturns. Identification of the impacts of such major events is always challenging since exogenous variation in exposure to historical shocks is rare. This paper may be unique in the timing of data collection in relation to the event in question: we are aware of no other study of risk preferences with data collection ongoing in the immediate run-up to and aftermath of a major event of this kind.

Existing studies of the impact of conflict on individual preferences estimate the marginal impact of greater exposure to violence, implicitly treating less-exposed survivors as a comparison group. We are able to expand this literature because our identification strategies enable us to measure the effect of civil conflict on the general population. Relatively more conflict-affected individuals may become more pro-social or less risk averse (Bellows and Miguel 2009, Blattman 2009, Voors, Nillesen, Verwimp, Bulte, Lensink, and Van Soest 2012, Callen, Isaqzadeh, Long, and Sprenger 2014), but the overall impact on the population as a whole is a shift toward less willingness to take profitable risks. Moreover, our results suggest that the channel of impact is social and not (primarily) economic, since the probability of engaging in any income-generating activity is not impacted by the violence, but generalized trust in others is. Our results therefore suggest that conflict may have long-lasting impacts on economic development through channels such as reduced entrepreneurship. That our findings differ from some earlier studies suggests that the impacts of conflict at the scale

of affected individuals and villages may not generalize to the scale of a district, or that of a nation.

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Table 1: Probit Regressions of Self-Employment and Migration on Risk Index

<i>Dependent Variable:</i>	SELF-EMPLOYED		MIGRATED SEEKING WORK	
	(1)	(2)	(3)	(4)
Times respondent chose riskiest lottery	0.006*** (0.002)	0.005** (0.003)	0.003** (0.001)	0.002* (0.001)
Demographic controls	No	Yes	No	Yes
Interviewer controls	No	Yes	No	Yes
Month controls	No	Yes	No	Yes
Observations	5049	5047	5047	5005
Pseudo R ²	0.002	0.095	0.004	0.195

Robust standard errors in parentheses. Marginal effects reported. Even-numbered columns include controls for gender, age, education level, marital status, and month and interviewer fixed effects. Independent variable is the number of times a respondent chose the riskiest lottery (out of five).

Table 2: OLS Regressions of the Impact of the Post-Election Crisis on Risk Preferences

<i>Specification:</i>	OLS (1)	OLS (2)	OLS (3)	OLS (4)	OLS (5)	OLS (6)	OLS (7)	OLS (8)
Surveyed after post-election violence	-0.683*** (0.116)	-0.686*** (0.113)	-0.798*** (0.118)	-0.79*** (0.116)	-0.561*** (0.084)	-0.571*** (0.09)	-0.735*** (0.086)	-0.734*** (0.088)
Demographic controls	No	Yes	No	Yes	No	Yes	No	Yes
Interviewer controls	No	No	Yes	Yes	No	No	Yes	Yes
Month controls	No	No	No	No	Yes	Yes	Yes	Yes
Observations	5049	5049	5047	5047	5049	5049	5047	5047
R^2	0.023	0.029	0.151	0.155	0.028	0.032	0.153	0.157

Robust standard errors clustered at the month level in all specifications. Outcome variable is the number of times a respondent chose the riskiest lottery (out of five). Even-numbered columns include controls for gender, age, education level, and marital status (at the time of the survey).

Table 3: Ordered Logit Regressions of the Impact of the Post-Election Crisis on Risk Preferences

<i>Specification:</i>	ORDERED LOGIT (1)	ORDERED LOGIT (2)	ORDERED LOGIT (3)	ORDERED LOGIT (4)	ORDERED LOGIT (5)	ORDERED LOGIT (6)	ORDERED LOGIT (7)	ORDERED LOGIT (8)
Surveyed after post-election violence	-0.641*** (0.104)	-0.641*** (0.101)	-0.787*** (0.12)	-0.773*** (0.116)	-0.54*** (0.088)	-0.546*** (0.097)	-0.749*** (0.098)	-0.742*** (0.102)
Demographic controls	No	Yes	No	Yes	No	Yes	No	Yes
Interviewer controls	No	No	Yes	Yes	No	No	Yes	Yes
Month controls	No	No	No	No	Yes	Yes	Yes	Yes
Observations	5049	5049	5047	5047	5049	5049	5047	5047
Pseudo R^2	0.007	0.009	0.045	0.047	0.008	0.01	0.046	0.048

Robust standard errors clustered at the month level in all specifications. Outcome variable is the number of times a respondent chose the riskiest lottery (out of five). Even-numbered columns include controls for gender, age, education level, and marital status (at the time of the survey).

Table 4: Regressions of the Impact of the Post-Election Crisis on Risk Preferences

<i>Specification:</i>	OLS (1)	OLS (2)	OLS (3)	OLS (4)	ORDERED LOGIT (5)	ORDERED LOGIT (6)	ORDERED LOGIT (7)	ORDERED LOGIT (8)
Surveyed after post-election violence	-0.607*** (0.184)	-0.686*** (0.113)	-0.359** (0.164)	-0.79*** (0.116)	-0.531*** (0.173)	-0.54*** (0.164)	-0.297* (0.171)	-0.315* (0.169)
Demographic controls	No	Yes	No	Yes	No	Yes	No	Yes
Interviewer controls	No	No	Yes	Yes	No	No	Yes	Yes
Linear time trends	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Observations	5049	5049	5047	5047	5049	5049	5047	5047
R^2	0.024	0.029	0.162	0.155
Pseudo R^2	0.007	0.01	0.05	0.052

Robust standard errors clustered at the month level in all specifications. Outcome variable is the number of times a respondent chose the riskiest lottery (out of five). Even-numbered columns include controls for gender, age, education level, and marital status (at the time of the survey).

Table 5: Robustness Checks: OLS Regressions of the Impact of the Post Election Violence on Alternate Measures of Risk Aversion

<i>Dependent Variable:</i>	RISK NEUTRAL OR LOVING (1)	MOST RISK AVERSE (2)	2 (3)	— CHOICE ON DECISION NUMBER... — 3 (4)	4 (5)	5 (6)	6 (7)
Surveyed after post-election violence	-0.057*** (0.017)	0.11*** (0.023)	-0.114*** (0.027)	-0.234*** (0.045)	-0.296*** (0.06)	-0.29*** (0.03)	-0.26*** (0.053)
Observations	5049	5049	5049	5049	5049	5049	5049
R^2	0.007	0.033	0.015	0.017	0.027	0.023	0.025

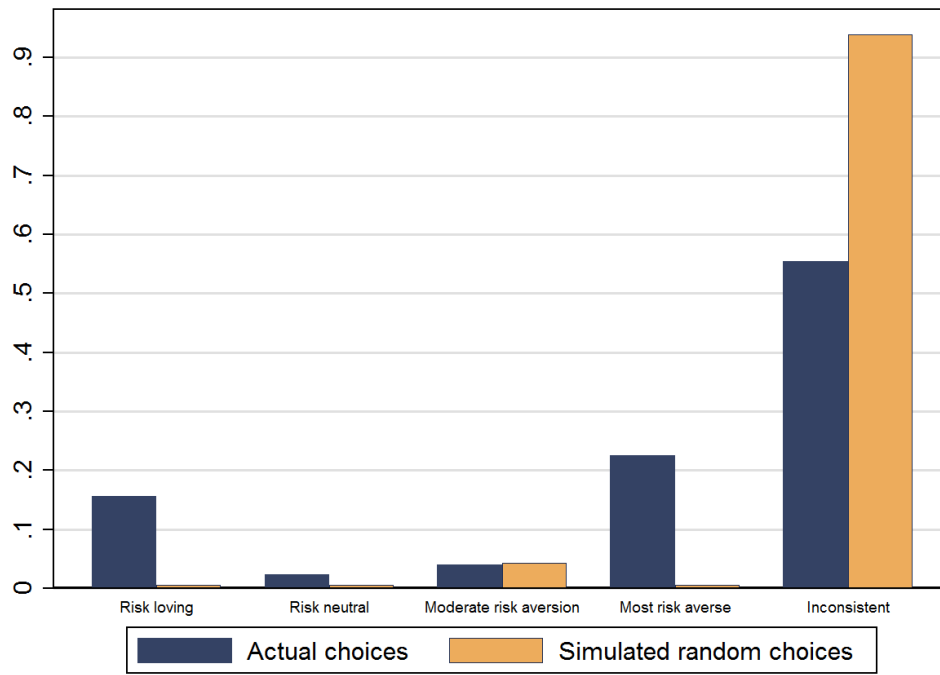
Robust standard errors clustered at the month level in all specifications. All specifications include controls for gender, age, education level, and marital status (at the time of the survey). Thus, the specifications in this table mirror those in Table 2, Column 2. Here, the outcome variables vary by column. The outcome in Column 1 is an indicator for making risk neutral or risk loving choices. Column 2 is an indicator for making the most risk averse choice on every card, so its sign is opposite those in other columns and tables. Columns 3 through 7 use discrete representations of the separate decision cards as outcomes, ordered by risk aversion; thus, the first two of these are indicator variables for taking the riskier option, while the remaining three are variables that take on the values 1, 2, and 3, with 1 being least risky, and 3 being riskiest.

Table 6: IV Regressions of the Impact of the Post-Election Crisis on Risk Preferences

	(1)	(2)	(3)	(4)
Surveyed after post-election violence	-1.050*** (0.223)	-1.779*** (0.5)	-1.085*** (0.257)	-1.988*** (0.588)
First Stage F-stat	9.210	8.050	13.660	11.560
Demographic controls	Yes	Yes	Yes	Yes
Interviewer controls	No	Yes	No	Yes
Month controls	No	No	Yes	Yes
Observations	5049	5047	5049	5047
R^2	0.022	0.118	0.024	0.117

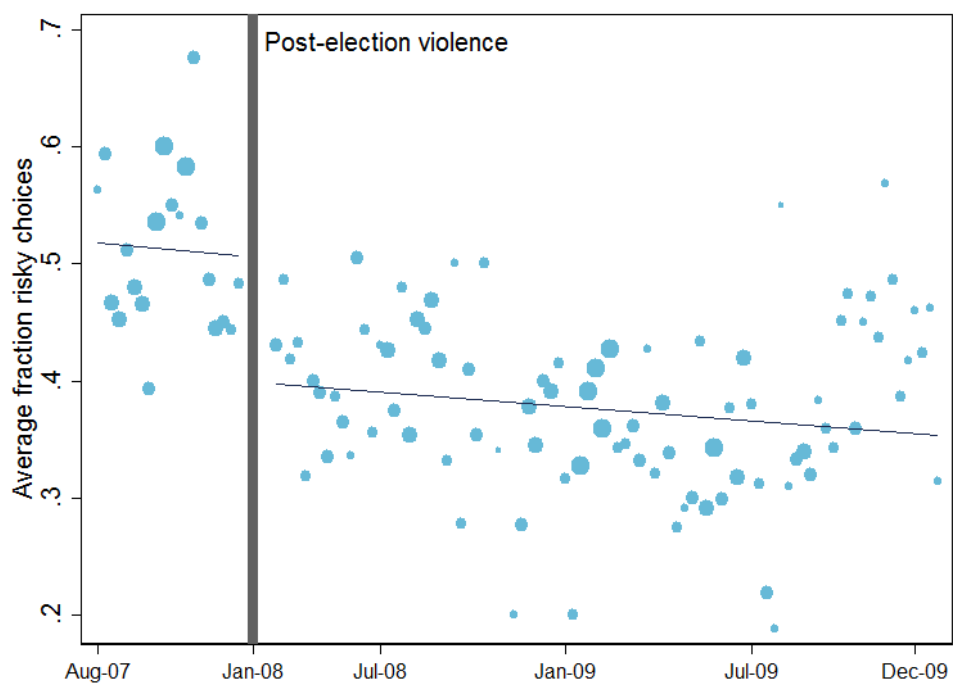
Robust standard errors clustered by survey month in all specifications. Outcome variable is the number of times a respondent chose the riskiest lottery (out of five).

Figure 1: Histogram of Patterns of Responses to Lottery Choice Questions



Dark bars show the proportion of subjects falling into each risk preference category. Lighter bars show the proportion that would fall into each category if subjects choose lotteries at random (i.e. if each lottery alternative were equally likely to be chosen in each decision problem).

Figure 2: Measured Risk Preferences Before and After Kenya's Post-Election Violence



Gray circles indicate the average number of risky choices by respondents surveyed in a given week, weighted by the number of surveys conducted during the week. Lines are linear regressions of the number of risk choices on the survey week, counting from the first week in which surveys were conducted.

A Online Appendix: not for print publication

Table A1: Decision Problems Embedded in Kenyan Life Panel Survey 2

	Option A		Option B		Option C	
	Heads	Tails	Heads	Tails	Heads	Tails
Practice Decision Problem 1	100	100	150	150	-	-
Practice Decision Problem 2	100	150	200	250	-	-
Decision Problem 1	100	100	100	120	-	-
Decision Problem 2	100	100	0	400	-	-
Decision Problem 3	30	340	100	100	0	400
Decision Problem 4	100	100	55	240	30	340
Decision Problem 5	30	230	60	170	90	110
Decision Problem 6	10	200	70	160	90	110

All payouts in Kenyan shillings.

Table A2: Balance Check: KLPS2 Wave1 vs. Wave 2 Comparison

Variable	WAVE 1	WAVE 2	DIFFERENCE
Number of observations	2469	2580	
Female	0.49 (0.01)	0.49 (0.01)	0.00 (0.01)
Age in 2007	21.13 (0.05)	21.21 (0.05)	0.09 (0.07)
Highest grade completed by 2006	8.56 (0.04)	8.48 (0.04)	-0.08 (0.06)
Married by 2006	0.31 (0.01)	0.30 (0.01)	-0.01 (0.01)

Standard errors in parentheses.

Table A3: Potential Mechanisms: Impacts on Labor Market Outcomes and Generalized Trust

	NON-FARM WORK (1)	WAGE WORK (2)	SELF- EMPLOYMENT (3)	TRUST OTHERS (4)
<i>Panel A: OLS</i>				
Surveyed after post-election violence	0.027 (0.029)	0.059*** (0.018)	-0.035** (0.016)	-0.057*** (0.011)
Observations	5049	5047	5049	5047
R^2	0.134	0.088	0.064	0.016
<i>Panel B: IV</i>				
Surveyed after post-election violence	-0.043 (0.06)	0.029 (0.053)	-0.071** (0.033)	-0.095*** (0.024)
First Stage F-stat	9.210	9.210	9.210	9.210
Observations	5049	5047	5049	5047
R^2	0.118	0.086	0.047	0.012

Robust standard errors clustered by survey month in all specifications. Panel A follows the specification in Table 2, Column 2. Panel B follows the specification in Table 6, Column 2. The outcomes in Columns 1 through 3 are indicator variables: the second and third columns are indicators for the respondent currently having wage work or self-employment; and the first column is an indicator for having either of these. In Column 4, the outcome is also binary, but is a representation of respondents' answer to the World Values Survey question: "Generally speaking, do you think that most people can be trusted or that you need to be very careful in dealing with people?"

Table A4: Robustness Checks: Impact of the Post Election Violence on Choosing Safest Alternative

<i>Dependent Variable:</i>	— SAFEST CHOICE ON DECISION NUMBER... —				
	2 (1)	3 (2)	4 (3)	5 (4)	6 (5)
Surveyed after post-election violence	0.114*** (0.027)	0.114*** (0.027)	0.11*** (0.029)	0.204*** (0.015)	0.18*** (0.027)
Observations	5049	5049	5049	5049	5049
R^2	0.015	0.017	0.019	0.037	0.033

Robust standard errors clustered at the month level in all specifications. All specifications include controls for gender, age, education level, and marital status (at the time of the survey). Thus, the specifications in this table mirror those in Table 2, Column 2. The outcome in each case is an indicator variable for taking the safest option in a given decision.

Figure A1: Histograms of Individual Choices in Each Decision Problem

